# Threshold Decoding for Disjunctive Group Testing ${ }^{1}$ 

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#### Abstract

Let $1 \leq s<t, N \geq 1$ be integers and a complex electronic circuit of size $t$ is said to be an $s$-active, $s \ll t$, and can work as a system block if not more than $s$ elements of the circuit are defective. Otherwise, the circuit is said to be an $s$-defective and should be replaced by a similar $s$-active circuit. Suppose that there exists a possibility to run $N$ non-adaptive group tests to check the $s$ activity of the circuit. As usual, we say that a (disjunctive) group test yields the positive response if the group contains at least one defective element. Along with the conventional decoding algorithm based on disjunctive $s$-codes, we consider a threshold decision rule with the minimal possible decoding complexity, which is based on the simple comparison of a fixed threshold $T, 1 \leq T \leq N-1$, with the number of positive responses $p, 0 \leq p \leq N$. For the both of decoding algorithms we discuss upper bounds on the $\alpha$-level of significance of the statistical test for the null hypothesis $\left\{H_{0}\right.$ : the circuit is $s$-active $\}$ verse the alternative hypothesis $\left\{H_{1}\right.$ : the circuit is $s$-defective $\}$.


## 1 Statement of Problem

Let $N \geq 2, t \geq 2, s$ and $T$ be integers, where $1 \leq s<t$ and $1 \leq T<N$. The symbol $\triangleq$ denote the equality by definition, $|A|-$ the size of the set $A$ and $[N] \triangleq\{1,2, \ldots, N\}$ - the set of integers from 1 to $N$. A binary $(N \times t)$-matrix

$$
\begin{equation*}
X=\left\|x_{i}(j)\right\|, \quad x_{i}(j)=0,1, \quad i \in[N], j \in[t] \tag{1}
\end{equation*}
$$

with $t$ columns (codewords) $\boldsymbol{x}(j) \triangleq\left(x_{1}(j), x_{2}(j), \ldots, x_{N}(j), j \in[t]\right.$, and $N$ rows $\boldsymbol{x}_{i} \triangleq\left(x_{i}(1), x_{i}(2) \ldots, x_{i}(t)\right), i \in[N]$, is called a binary code of length $N$ and size $t=\left\lfloor 2^{R N}\right\rfloor$, where a fixed parameter $R>0$ is called a rate of the code $X$. The number of 1's in a binary column $\boldsymbol{x}=\left(x_{1}, \ldots, x_{N}\right) \in\{0,1\}^{N}$, i.e., $|\boldsymbol{x}| \triangleq \sum_{i=1}^{N} x_{i}$, is called a weight of $\boldsymbol{x}$. A code $X$ is called a constant weight binary code of weight $w, 1 \leq w<N$, if for any $j \in[t]$, the weight $|\boldsymbol{x}(j)|=w$. The conventional symbol $\boldsymbol{u} \bigvee \boldsymbol{v}$ will be used to denote the disjunctive (Boolean)

[^0]sum of binary columns $\boldsymbol{u}, \boldsymbol{v} \in\{0,1\}^{N}$. We say that a column $\boldsymbol{u}$ covers a column $\boldsymbol{v}(\boldsymbol{u} \succeq \boldsymbol{v})$ if $\boldsymbol{u} \bigvee \boldsymbol{v}=\boldsymbol{u}$.

### 1.1 Disjunctive and Threshold Disjunctive Codes

Definition 1. [1]. A code $X(1)$ is called a disjunctive $s$-code, $s \in[t-1]$, if the disjunctive sum of any $s$-subset of codewords of $X$ covers those and only those codewords of $X$ which are the terms of the given disjunctive sum.

Let $\mathcal{S}, \mathcal{S} \subset[t]$, be an arbitrary fixed collection of defective elements of size $|\mathcal{S}|$. For a binary code $X$ and collection $\mathcal{S}$, define the binary response vector of length $N$, namely:

$$
\begin{equation*}
x(\mathcal{S}) \triangleq \bigvee_{j \in \mathcal{S}} x(j), \quad \text { if } \quad \mathcal{S} \neq \emptyset \quad \text { and } \quad x(\mathcal{S}) \triangleq(0,0, \ldots, 0) \quad \text { if } \quad \mathcal{S}=\emptyset \tag{2}
\end{equation*}
$$

In the classical problem of non-adaptive group testing, we describe $N$ tests as a binary $(N \times t)$-matrix $X=\left\|x_{i}(j)\right\|$, where a column $\boldsymbol{x}(j)$ corresponds to the $j$-th element, a row $\boldsymbol{x}_{i}$ corresponds to the $i$-th test and $x_{i}(j) \triangleq 1$ if and only if the $j$-th element is included into the $i$-th testing group. The result of each test equals 1 if at least one defective element is included into the testing group and 0 otherwise, so the column of results is exactly equal to the response vector $\boldsymbol{x}(\mathcal{S})$. Definition 1 of disjunctive $s$-code $X$ gives the important sufficient condition for the evident identification of any unknown collection of defective elements $\mathcal{S}$ if the number of defective elements $|\mathcal{S}| \leq s$. In this case, the identification of the unknown $\mathcal{S}$ is equivalent to discovery of all codewords of code $X$ covered by $\boldsymbol{x}(\mathcal{S})$, and its complexity is equal to the code size $t$. Note that this algorithm also allows us to check $s$-activity of the circuit defined in the abstract. Moreover, it is easy to prove by contradiction that every code $X$ which allows to check $s$-activity of the circuit without error is a disjunctive $s$-code.

Definition 2. Let $s, s \in[t-1]$, and $T, T \in[N-1]$, be arbitrary fixed integers. A disjunctive $s$-code $X$ of length $N$ and size $t$ is said to be a a disjunctive $s$-code with threshold $T$ (or, briefly, $s^{T}$-code) if the disjunctive sum of any $\leq s$ codewords of $X$ has weight $\leq T$ and the disjunctive sum of any $\geq s+1$ codewords of $X$ has weight $\geq T+1$.

Obviously, for any $s$ and $T$, the definition of $s^{T}$-code gives a sufficient condition for code $X$ applied to the group testing problem described in the abstract of our paper. In this case, only on the base of the known number of positive responses $|\boldsymbol{x}(\mathcal{S})|$, we decide that the controllable circuit identified by an unknown collection $\mathcal{S}, \mathcal{S} \subset[t]$, is $s$-active, i.e., the unknown size $|\mathcal{S}| \leq s$ ( $s$-defective, i.e., the unknown size $|\mathcal{S}| \geq s+1)$ if $|\boldsymbol{x}(\mathcal{S})| \leq T(|\boldsymbol{x}(\mathcal{S})| \geq T+1)$.

Remark 1. The concept of $s^{T}$-codes was motivated by troubleshooting in complex electronic circuits using a non-adaptive identification scheme which was considered in [2].

### 1.2 Hypothesis Test

Let a circuit of size $t$ is identified by an unknown collection $\mathcal{S}_{u n}, \mathcal{S}_{u n} \subset[t]$, of defective elements of an unknown size $\left|\mathcal{S}_{u n}\right|$ and $X$ be a code (1) of size $t$ and length $N$. Introduce the null hypothesis $\left\{H_{0}:\left|\mathcal{S}_{u n}\right| \leq s\right\}$ (the circuit is $s$-active) verse the alternative $\left\{H_{1}:\left|\mathcal{S}_{u n}\right| \geq s+1\right\}$ (the circuit is $s$-defective). In this paper we focus on the testing of statistical hypotheses $H_{0}$ and $H_{1}$. The similar problem related to constructing of a confidence interval for $\left|\mathcal{S}_{u n}\right|$ was considered in [3], where the authors construct the interval $[\hat{s} / c ; \hat{s}]$, such that given a random code $X$, the statistic $\hat{s}$, i.e., a function of the random response vector $\boldsymbol{x}\left(\mathcal{S}_{u n}\right)$, satisfies the following properties: $\operatorname{Pr}\left\{\hat{s}<\left|\mathcal{S}_{u n}\right|\right\}$ is upper bounded by a small parameter $\epsilon \ll 1$ and the expected value of $\hat{s} /\left|\mathcal{S}_{u n}\right|$ is upper bounded by a number $c>1$.

For fixed parameters $s, 1 \leq s<t$, and $T, 1 \leq T<N$, consider the following threshold decision rule motivated by Definition 2, namely:

$$
\begin{cases}\text { accept }\left\{H_{0}:\left|\mathcal{S}_{u n}\right| \leq s\right\} & \text { if }\left|\boldsymbol{x}\left(\mathcal{S}_{u n}\right)\right| \leq T,  \tag{3}\\ \text { accept }\left\{H_{1}:\left|\mathcal{S}_{u n}\right| \geq s+1\right\} & \text { if }\left|\boldsymbol{x}\left(\mathcal{S}_{u n}\right)\right| \geq T+1\end{cases}
$$

For the conventional statistical interpretation of the decision rule (3), it is reasonable to assume that the different collections of defective elements of the same size are equiprobable. That is why, we set that the probability distribution of the random collection $\mathcal{S}_{u n}, \mathcal{S}_{u n} \subset[t]$, is identified by an unknown probability vector $\boldsymbol{p} \triangleq\left(p_{0}, p_{1}, \ldots, p_{t}\right), p_{k} \geq 0, k=0,1, \ldots, t, \sum_{k=0}^{t} p_{k}=1$, as follows:

$$
\begin{equation*}
\operatorname{Pr}\left\{\mathcal{S}_{u n}=\mathcal{S}\right\} \triangleq \frac{p_{|\mathcal{S}|}}{\left(\left|{ }_{|S|}\right|\right.} \quad \text { for any subset } \quad \mathcal{S} \subseteq[t] \tag{4}
\end{equation*}
$$

Introduce a maximal error probability of the decision rule (3) :

$$
\begin{equation*}
\varepsilon_{s}(T, \boldsymbol{p}, X) \triangleq \max \left\{\operatorname{Pr}\left\{\text { accept } H_{1} \mid H_{0}\right\}, \operatorname{Pr}\left\{\text { accept } H_{0} \mid H_{1}\right\}\right\} \tag{5}
\end{equation*}
$$

where the conditional probabilities in the right-hand side of (5) are identified by (3)-(4). Note that the number $\varepsilon_{s}(T, \boldsymbol{p}, X)=0$ if and only if the code $X$ is an $s^{T}$-code. Denote by $t_{s}(N, T)$ the maximal size of $s^{T}$-codes of length $N$. For a parameter $\tau, 0<\tau<1$, introduce the rate of $s^{\lfloor\tau N\rfloor}$-codes:

$$
\begin{equation*}
R_{s}(\tau) \triangleq \varlimsup_{N \rightarrow \infty} \frac{\log _{2} t_{s}(N,\lfloor\tau N\rfloor)}{N} \geq 0, \quad 0<\tau<1 \tag{6}
\end{equation*}
$$

Definition 3. Let $\tau, 0<\tau<1$, and a parameter $R, R>R_{s}(\tau)$, be fixed. For the maximal error probability $\varepsilon_{s}(T, \boldsymbol{p}, X)$, defined by (3)-(5), consider the function

$$
\begin{equation*}
\varepsilon_{s}^{N}(\tau, R) \triangleq \max _{p}\left\{\min _{X} \varepsilon_{s}(\lfloor\tau N\rfloor, \boldsymbol{p}, X)\right\}, \quad R>R_{s}(\tau) \tag{7}
\end{equation*}
$$

where the minimum is taken over all codes $X$ of length $N$ and size $t=\left\lfloor 2^{R N}\right\rfloor$ with parameter $R>R_{s}(\tau)$. The number $\varepsilon_{s}^{N}(\tau, R)>0$ does not depend on the unknown probability vector $\boldsymbol{p}$ and can be called the universal error probability of the decision rule (3). The corresponding error exponent

$$
\begin{equation*}
E_{s}(\tau, R) \triangleq \varlimsup_{N \rightarrow \infty} \frac{-\log _{2} \varepsilon_{s}^{N}(\tau, R)}{N}, \quad s \geq 1, \quad 0<\tau<1, \quad R>R_{s}(\tau) \tag{8}
\end{equation*}
$$

identifies the asymptotic behavior of $\alpha$-level of significance for the decision rule (3), i.e.,

$$
\begin{equation*}
\alpha \triangleq \exp _{2}\left\{-N\left[E_{s}(\tau, R)+o(1)\right]\right\}, \quad \text { if } \quad E_{s}(\tau, R)>0, \quad N \rightarrow \infty \tag{9}
\end{equation*}
$$

Along with (3) we introduce the disjunctive decision rule based on the conventional algorithm:

$$
\begin{cases}\text { accept } H_{0} & \text { if } \boldsymbol{x}\left(\mathcal{S}_{u n}\right) \text { covers } \leq s \text { codewords of } X  \tag{10}\\ \text { accept } H_{1} & \text { if } \boldsymbol{x}\left(\mathcal{S}_{u n}\right) \text { covers } \geq s+1 \text { codewords of } X .\end{cases}
$$

For a fixed code rate $R, R>0$, the error exponent for disjunctive decision rule (10) $E_{s}(R)$ is defined similarly to (5)-(8). The function $E_{s}(R)$ was firstly introduced in our paper [1], where we proved

Theorem 1. [1]. The function $E_{s}(R)=0$ if $R \geq 1 / s$.

## 2 Lower Bounds on Error Exponents

In this Section, we formulate and compare the random coding lower bounds for the both of error exponents $E_{s}(R)$ and $E_{s}(\tau, R)$. These bounds were proved applying the random coding method based on the ensemble of constant-weight codes. A parameter $Q$ in formulations of theorems 2-3 means the relative weight of codewords of constant-weight codes. Introduce the standard notations

$$
h(Q) \triangleq-Q \log _{2} Q-(1-Q) \log _{2}[1-Q], \quad[x]^{+} \triangleq \max \{x, 0\} .
$$

In [1], we established
Theorem 2. [1]. The error exponent $E_{s}(R) \geq \underline{E}_{s}(R)$ where the random coding lower bound

$$
\begin{aligned}
& \underline{E}_{s}(R) \triangleq \max _{0<Q<1} \min _{Q \leq q<\min \{1, s Q\}}\left\{\mathcal{A}(s, Q, q)+[h(Q)-q h(Q / q)-R]^{+}\right\}, \\
& \mathcal{A}(s, Q, q) \triangleq(1-q) \log _{2}(1-q)+q \log _{2}\left[\frac{Q y^{s}}{1-y}\right]+s Q \log _{2} \frac{1-y}{y}+s h(Q),
\end{aligned}
$$

and $y$ is the unique root of the equation

$$
q=Q \frac{1-y^{s}}{1-y}, \quad 0<y<1 .
$$

In addition, as $s \rightarrow \infty$ and $R \leq \frac{\ln 2}{s}(1+o(1))$, the lower bound $\underline{E}_{s}(R)>0$.
Theorem 3. 1. The error exponent $E_{s}(\tau, R) \geq \underline{E}_{s}(\tau, R)$ where the random coding bound $\underline{E}_{s}(\tau, R)$ does not depend on $R>0$ and has the form:

$$
\begin{gathered}
\underline{E}_{s}(\tau, R) \triangleq \max _{1-(1-\tau)^{1 /(s+1)<Q<1-(1-\tau)^{1 / s}} \min \left\{\mathcal{A}^{\prime}(s, Q, \tau), \mathcal{A}(s+1, Q, \tau)\right\},} \begin{array}{l}
\mathcal{A}^{\prime}(s, Q, q) \triangleq \begin{cases}\mathcal{A}(s, Q, q), & \text { if } Q \leq q \leq s Q \\
\infty, & \text { otherwise. }\end{cases}
\end{array} .\left\{\begin{array}{l}
\text {, }
\end{array}\right.
\end{gathered}
$$

2. As $s \rightarrow \infty$ the optimal value of $\underline{E}_{s}(\tau, R)$ satisfies the inequality:

$$
\underline{E}_{T h r}(s) \triangleq \max _{0<\tau<1} \underline{E}_{s}(\tau, R) \geq \frac{\log _{2} e}{4 s^{2}}(1+o(1)), \quad s \rightarrow \infty
$$

It is possible to use decision rule (3) with any value of parameter $T$. The numerical values of the optimal error exponent $\underline{E}_{\text {Thr }}(s)$ along with the corresponding optimal values of threshold parameter $\tau$ and the constant-weight code ensemble parameter $Q$ are presented in Table 1. Besides in Table 1 the values of $\underline{E}_{s}(0)$ and $R_{\text {cr }} \triangleq \sup \left\{R: \underline{E}_{s}(R)>\underline{E}_{\mathrm{Thr}}(s)\right\}$ are shown. Theorems 1-3 show that, for large values of $R$, the threshold decision rule (3) has an advantage over the disjunctive decision rule (10) as $N \rightarrow \infty$.

Table 1: The numerical values of $\underline{E}_{\text {Thr }}(s)$

| $s$ | 2 | 3 | 4 | 5 | 6 | 7 | 8 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\underline{E}_{\mathrm{Thr}}(s)$ | 0.1380 | 0.0570 | 0.0311 | 0.0196 | 0.0135 | 0.0098 | 0.0075 |
| $\tau$ | 0.2065 | 0.1365 | 0.1021 | 0.0816 | 0.0679 | 0.0582 | 0.0509 |
| $Q$ | 0.1033 | 0.0455 | 0.0255 | 0.0163 | 0.0113 | 0.0083 | 0.0064 |
| $\underline{E}_{s}(0)$ | 0.3651 | 0.2362 | 0.1754 | 0.1397 | 0.1161 | 0.0994 | 0.0869 |
| $R_{\text {cr }}$ | 0.2271 | 0.1792 | 0.1443 | 0.1201 | 0.1027 | 0.0896 | 0.0794 |

## 3 Simulation for finite code parameters

For finite $N$ and $t$, we carried out a simulation as follows. The probability distribution vector $\boldsymbol{p}$ (4) is defined by

$$
p_{k} \triangleq\binom{t}{k} p^{k}(1-p)^{t-k}, \quad p \triangleq \frac{s+1 / 2}{t}, \quad 0 \leq k \leq t
$$

i.e. the number of defective elements $\left|\mathcal{S}_{u n}\right|$ is binomially distributed and has the expected value $s+1 / 2$. A code $X$ is generated randomly from the ensemble of constant-weight codes, i.e. for some weight parameter $w$, every codeword of $X$ is chosen independently and equalprobably from the set of all $\binom{t}{w}$ codewords. For every weight $w$ and every decision rule, we repeat generation 1000 times and choose the code with minimal error probability. Note that for disjunctive decision rule $\operatorname{Pr}\left\{\right.$ accept $\left.H_{0} \mid H_{1}\right\}=0$. In Table 2 the best values of maximal error probability for fixed parameters $s, t$ and $N$ are shown in bold.

Table 2: Results of simulation

|  | Threshold decision rule |  |  |  | Disjunctive decision rule |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $N$ | $\operatorname{Pr}\left\{\right.$ acc. $\left.H_{1} \mid H_{0}\right\}$ | $\operatorname{Pr}\left\{\right.$ acc. $\left.H_{0} \mid H_{1}\right\}$ | $w$ | $T$ | $\operatorname{Pr}\left\{\right.$ acc. $\left.H_{1} \mid H_{0}\right\}$ | $w$ |
| s=2, $t=15$ |  |  |  |  |  |  |
| 5 | 0.1366 | 0.1355 | 2 | 3 | 0.4780 | 2 |
| 8 | 0.0732 | 0.0824 | 3 | 5 | 0.3610 | 2 |
| 10 | 0 | 0.0744 | 1 | 2 | 0.2390 | 3 |
| 12 | 0 | 0.0440 | 1 | 2 | 0.1220 | 3 |
| 14 | 0 | 0.0349 | 2 | 4 | 0.0537 | 3 |
| 15 | 0 | 0.0258 | 2 | 4 | 0.0195 | 3 |
| $s=2, \quad t=20$ |  |  |  |  |  |  |
| 5 | 0.1398 | 0.1365 | 2 | 3 | 0.5356 | 2 |
| 8 | 0.0897 | 0.0890 |  | 5 | 0.4169 | 2 |
| 10 | 0.0897 | 0.0858 | 3 | 5 | 0.3008 | 3 |
| 12 | 0.0580 | 0.0576 | 4 | 7 | 0.1979 | 3 |
| 15 | 0 | 0.0324 | 2 | 4 | 0.0792 | 4 |

## References

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